

Impact of dental implant length on early failure rates: a metaanalysis of observational studies

Pommer B, Frantal S, Willer J, Posch M, Watzek G, Tepper G. Impact of dental implant length on early failure rates: a meta-analysis of observational studies. J Clin Periodontol 2011; 38: 856–863. doi: 10.1111/j.1600-051X.2011.01750.x.

Abstract

Aim: To test the null hypothesis of no difference in failure rates of short (minimum length: 7 mm) and longer dental implants (≥ 10 mm), a meta-analysis was performed on prospective observational trials.

Materials and Methods: A systematic electronic and hand search was performed to identify eligible studies. Having additional data supplied by the authors, 54 publications were included (19,083 implants).

Results: In case of mandibular implants, the null hypothesis of no impact of reduced implant length on failure within the first year of prosthetic loading could not be rejected. A significant impact of implant length could be substantiated for short machined implants in the anterior [odds ratio (OR) 5.4] and posterior maxilla (OR 3.4), while short rough-surfaced implants demonstrated increased failure rates in the anterior maxillary sites. No influence of implant diameter and denture type on the failure rate of short implants could be revealed.

Conclusion: In areas of reduced alveolar bone height the use of short dental implants may reduce the need for invasive bone augmentation procedures.

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Key words: diameter; failure rate; length; short implants; success rate

Accepted for publication 21 May 2011

Endosseous dental implants to replace the natural tooth provide a reliable basis for fixed and removable dentures (Stellingsma et al. 2004). Reduced alveolar bone height due to post-extraction ridge resorption and maxillary sinus pneumatization represents a major limitation in the use of dental implants (Tawil & Younan 2003). Surgical treatment options to overcome this limitation comprise either supplementary bone augmentation procedures or the exclusive use of implants reduced in length (das Neves et al. 2006). The key advantage of placing short implants is the avoidance of invasive bone augmentation

Conflict of interest and source of funding statement

The authors declare that there are no conflicts of interests. No external funding was obtained. surgery associated with donor site morbidity, additional treatment duration and financial burden (Nedir et al. 2004). Further benefits include a reduced risk of sinus perforation and mandibular paraesthesia, as well as the possibility to obviate pre-surgical diagnostic radiography (Misch et al. 2006, Morand & Irinakis 2007). Positive clinical results have increased the interest in this promising technique (Romeo et al. 2006, Tawil et al. 2006, Maló et al. 2007). However, increased failure rates have been observed when placing dental implants under 10 mm in length (Bahat 2000, Attard & Zarb 2003, Weng et al. 2003), indicating a tendency of failures to occur either within the healing phase or the first year of prosthetic loading.

Current literature is still controversial in regard to the reliability and indications of short dental implants. Systematic reviews did not reveal a clear correlation between implant length and implant success rates and evidence from randomized controlled trials (RCTs) is still missing (Hagi et al. 2004, Stellingsma et al. 2004, das Neves et al. 2006, Renouard & Nisand 2006, Kotsovilis et al. 2009). Meta-analysis of RCTs is the gold standard for appraising evidence from clinical trials (Sterne et al. 2001); however, meta-analyses of nonrandomized studies allow interventions to be evaluated that are not feasible to investigate by RCTs (Deeks et al. 2003, Reeves & Gaus 2004). As it seems unethical to investigate short implants in jaw regions of large bone volumes in RCTs, it was decided to perform a metaanalysis of prospective observational studies. The aim of the present investigation was to test the null hypothesis of no impact of reduced implant length on implant failure within the first year of loading.

Materials and Methods

Literature search and selection

A search of English literature published from January, 1998, to January, 2008, was conducted involving an electronic MEDLINE, EMBASE, and CENTRAL search and hand searching of 29 journals (supporting information Table S1). Studies were considered if they met following inclusion criteria: (1) prospective clinical investigations, defined as observational studies where allocation occurs in the course of usual treatment decisions (Higgings & Green 2008), reporting implant failure rates within the first year of prosthetic loading based on established criteria (Albrektsson et al. 1986), (2) endosseous titanium implants placed in non-augmented, healed jawbone and loaded after a conventional healing period of 3-6 months (Degidi et al. 2006), and (3) implants shorter than 10 mm (minimum length: 7 mm) included. Trials investigating patients with general systemic illness or under 20 years of age were not considered. Two reviewers (B. P. and G. T.) independently screened the titles and abstracts of the search results. Full texts of all papers that were considered eligible for inclusion by one or both of the reviewers were obtained for further assessment against the stated inclusion criteria. Disagreements were resolved by consensus. When multiple reports on the same patients were identified, the most recent publication was included.

Data abstraction and collection

Out of 1363 articles screened, 119 publications were selected as preliminary candidates and underwent data abstraction in duplicate. The Newcastle-Ottawa scale (NOS) was used to assess study quality (Wells et al. 2001). Studies that received NOS ratings of ≥ 7 stars (of nine possible stars) were judged as "high quality" (Chak et al. 2009). For each trial following information was recorded: implant system and surface texture, number of implants lost to follow-up, and type of prosthetic rehabilitation. The implant data was subdivided according to implant dimension and jaw position; however, the majority of studies did not detail individual failure scores for the various implant lengths and diameters. Owing to inappropriate presentation or limited information provided in the articles, breakdown of data corresponding to the first year of obser-



Fig. 1. Flow chart for the search process indicating numbers (N) of excluded studies, stages of exclusion, and reasons for exclusion.

vation could rarely be achieved. Hence a total of 95 authors were contacted for clarification or missing data (119 articles). The corresponding authors received prepared data forms by postal mail, fax, and e-mail and were sent a reminder after 3 and 6 months. Studies were not considered if the required information was not obtained within an editing time of 8 months. After taking into account the additional data kindly provided by the authors, 54 trials constituted the final selection (Fig. 1).

Quantitative data synthesis

Outcome measure was implant failure within the first year of prosthetic loading defined as implant mobility, infection, pain, peri-implant radiolucency, or progressive marginal bone loss (Albrektsson et al. 1986). Implants in patients lost to follow-up were not analysed. Two types of analyses were performed: (1) within-study comparison: To estimate the effect of implant length, diameter, and surface texture on failure rates. Mantel-Haenszel tests for the overall odds-ratios, Fischer-exact-tests for the odds-ratios in each study (including 95% confidence intervals), and Woolf's tests for statistical heterogeneity were performed. To correct for multiple testing (for the three overall Mantel-Haenszel tests) a Bonferroni corrected two-side significance level of 0.05/ 3 = 0.0167 was applied. The impact of implant length was separately tested for machined- as well as rough-surfaced implants for following subgroups: anterior maxilla, posterior maxilla, anterior mandible, and posterior mandible. In each of the above analyses, only studies with at least 10 implants in each group were included. (2) Descriptive analyses: From the data pooled across all 54 studies failure rates and 95% confidence intervals (based on the binomial test) were computed according to implant diameter and type of denture. All analyses were performed using R 2.4.0 (R Foundation for Statistical Computing, Vienna, Austria).

Results

The mean NOS score of the 54 included trials was 7.2 \pm 0.5 and ranged between 7 and 9 (supporting information Table S2) indicating high methodological quality throughout the included studies (Table 1). A total of 19,563 implants were placed, of which 480 implants could not be analysed due to patient dropout within the first year (mean drop-out rate: 2.0%). Of the 19,083 implants included, 40.1% were placed in the maxilla and 59.9% were placed in the mandible. 49.3% of the implants were inserted in incisor and canine regions (anterior positions), the other half in pre-molar and molar regions (posterior positions). According to their surface texture, the implants were grouped into 8686 machined and 10,397 rough implants (Khang et al. 2001). 1880 implants supported fixed restorations (single crowns or fixed partial dentures), and 6865 implants supported removable dentures. The mean implant

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| Table 1. Descriptive information on the 54 include | ed studies |
|--|------------|
|--|------------|

| Study | Total number of implants | Number of implant failures | Number of implants lost to follow-up | Implant manufacturer | Implant surface | Type of denture |
|--------------------------------|-----------------------------|----------------------------|--------------------------------------|-------------------------|-----------------|-----------------|
| Amileon et al. (1008) | (19 | 7 | 0 | Δ | Mashinad | Eined |
| Astrand at al. (2000) | 167 | 12 | 9 | Astra Tech | Pough | Fixed |
| Astrand et al. (2000) | 107 | 12 | 0 | Various | Doth | Fixed |
| Astrand at al. $(2004a)$ | 571 | 9 | 0 | Various | Doth | Fixed |
| Astrand et al. (2004b) | 150 | 3 19 | 0 | Various | Both | Fixed |
| Attard and Zarb (2003) | 398 | 18 | 5 25 | Nobel Biocare | Machined | Fixed |
| Banat (2000) | 652 | 25 | 35 | Nobel Biocare | Machined | Fixed |
| Bakke et al. (2002) | 24 | 0 | 0 | Astra Tech | Machined | Removable |
| Balleri et al. (2002) | 45 | 0 | 0 | Nobel Biocare | Machined | Fixed |
| Behneke et al. (2000) | 114 | 0 | 0 | Straumann | Rough | Fixed |
| Behneke et al. (2002) | 340 | 4 | 1 | Straumann | Rough | Removable |
| Bergkvist et al. (2004) | 144 | 5 | 0 | Straumann | Rough | Fixed |
| Bischof et al. (2004) | 43 | 1 | 0 | Straumann | Rough | Fixed |
| Bischof et al. (2006) | 259 | 2 | 4 | Straumann | Rough | Fixed |
| Brocard et al. (2000) | 830 | 11 | 0 | Straumann | Rough | Both |
| Degidi et al. (2006) | 521 | 3 | 0 | Friatec | Rough | Fixed |
| Deporter et al. (1998) | 20 | 0 | 0 | Endopore | Rough | Fixed |
| Deporter et al. (1999) | 156 | 6 | 0 | Endopore | Rough | Removable |
| Deporter et al. (2001a) | 48 | 0 | 0 | Endopore | Rough | Fixed |
| Deporter et al. (2001b) | 149 | 4 | 0 | Endopore | Rough | Fixed |
| Eliasson et al. (2000) | 476 | 2 | 0 | Nobel Biocare | Machined | Fixed |
| Ferrigno et al. (2002) | 1044 | 4 | 0 | Straumann | Rough | Both |
| Friberg et al. (2000) | 247 | 6 | 4 | Nobel Biocare | Machined | Both |
| Friberg et al. (2003) | 88 | 8 | 0 | Nobel Biocare | Machined | Fixed |
| Friberg et al. (2005) | 451 | 5 | 60 | Nobel Biocare | Rough | Both |
| Gaucher et al. (2001) | 688 | 5 | 3 | Implant Innovations | Rough | Both |
| Gotfredsen and Karlsson (2001) | 128 | 2 | 18 | Astra Tech | Both | Fixed |
| Grunder et al. (1999) | 219 | 3 | 14 | Implant Innovations | Rough | Both |
| Hallman (2001) | 182 | 1 | 1 | Straumann | Rough | Both |
| Jemt and Johansson (2006) | 450 | 9 | 20 | Nobel Biocare | Machined | Fixed |
| Khang et al. (2001) | 432 | 36 | 20 | Implant Innovations | Both | Both |
| Künzel et al. (2002) | 432 | 1 | 12 | Straumann | Rough | Both |
| Mattsson et al. (1999) | 86 | 1 | 0 | Nobel Biocare | Machined | Fixed |
| Moheng and Feryn (2005) | 266 | 9 | 0 | Various | Rough | Both |
| Nedir et al. (2006) | 522 | 2 | 25 | Straumann | Rough | Both |
| Ortorp and Jemt (2004) | 367 | 0 | 11 | Nobel Biocare | Machined | Removable |
| Palmer et al. (2005) | 21 | 0 | 2 | Astra Tech | Rough | Fixed |
| Preiskel and Tsolka (2004) | 269 | 5 | 2 | Nobel Biocare | Machined | Fixed |
| Renouard et al. (1999) | 59 | 4 | 0 | Nobel Biocare | Machined | Fixed |
| Romeo et al. (2002) | 187 | 0 | 4 | Straumann | Rough | Fixed |
| Romeo et al. (2003) | 100 | 2 | 0 | Various | Both | Fixed |
| Romeo et al. (2004) | 759 | -2 | Ő | Straumann | Rough | Both |
| Romeo et al. (2006) | 265 | 0 | Ő | Straumann | Rough | Both |
| Sethi et al. (2000) | 205 | 33 | 141 | Various | Both | Both |
| Tangerud et al. (2000) | 85 | 5 | 0 | Nobel Biocare | Machined | Fixed |
| Tawil and Younan (2003) | 253 | 4 | 9 | Nobel Biocare | Machined | Fixed |
| Tawil at al. (2006) | 233 | 0 | 0 | Nobel Biocare | Machined | Fixed |
| Testori et al. (2000) | 485 | 6 | 34 | Implant Innovations | Pough | Fixed |
| Viscer et al. (2001) | 106 | 8 | 0 | Various | Rough | Perrovable |
| Weng et al. (2000) | 190 874 | 0 ∕[1 | 0 | Implant Inpovations | Machinad | Fived |
| Wonnerbarg at a^{1} (2003) | 520 | +1 1 | 7 | Nobal Diagona | Machined | Domovahl- |
| Widmark at al. (2001) | J38 117 | 1 | 4 | Nobel Biocare | Machined | Removable |
| Widmark et al. (2001) | 11/ | 5 | 9 | Nobel Biocare | Machined | BOIN D-41- |
| within the tail (2003) | 194 | 5 11 | 10 | Nobel Blocare | Day -1- | BOIN D-41- |
| wheret al. (2003) | 1250 | 11 | 2 | Friatec | Rougn | BOIN |
| wismeijer et al. (1999) | 281 | 4 | 0 | Straumann | Rough | Removable |

Astra Tech AB (Mölndal, Sweden), Straumann AG (Waldenburg, Switzerland), Nobel Biocare (Gothenburg, Sweden), Friatech (Mannheim, Germany), Implant Innovations (West Palm Beach, FL, USA), Endopore (Innova Corporation, Toronto, ON, Canada).

diameter accounted 4.0 mm (range: 2.75–6.5 mm): 2568 implants were classified as narrow-diameter implants (<3.75 mm), and 16,515 implants as regular-diameter implants (≥ 3.75 mm). Implant length ranged between 7 and 20 mm: 2581 implants were shorter than

10 mm in length (short implants), while 16,502 implants were as least 10 mm long.

Within-study comparison of short and long implants was possible in 40 trials and is illustrated as a forest plot (Fig. 2). Short implants demonstrated a significantly higher overall failure rate compared with longer implants [odds ratio (OR) 1.8] with significant differences observed in the anterior (OR 6.1) and posterior (OR 3.6) maxilla, while no effect could be seen in the mandible (Table 2). Rough-surfaced implants showed significantly lower failure rates than machined ones (OR 3.6) and were therefore tested separately: While machined implants still demonstrated a significant impact of implant length overall as well as in the anterior and posterior maxilla (OR 2.2, 5.4, and 3.4, respectively), no significant overall effect could be substantiated for roughsurfaced implants (OR 1.1). In the anterior maxilla, short rough-surfaced implants demonstrated significantly increased failure rate (1.4% versus 0.0%), yet, no OR could be calculated due to the absence of long implant failures. In conclusion, the null hypothesis

| Study(Year) | ns | nl | OR[CI] | | |
|------------------|-----|------|-------------------|------------------------------|-----------------------------|
| Arvidson (1998) | 118 | 491 | 1.7 [0.2,10.4] | • • | |
| Deporter (1999) | 95 | 61 | 0.3 [0.0,2.2] | | |
| Grunder (1999) | 30 | 175 | 3.0 [0.0,58.6] | ~ · · · | ••••• |
| Renouard (1999) | 26 | 33 | 1.3 [0.1,18.9] | | |
| Wismeijer (1999) | 60 | 221 | 1.2 [0.0,15.6] | | |
| Bahat (2000) | 88 | 529 | 4.4 [1.7,10.8] | | - |
| Brocard (2000) | 198 | 632 | 0.7 [0.1,3.5] | <-∎ | - |
| Sethi (2000) | 80 | 2040 | 1.7 [0.2,6.7] | | |
| Deporter1 (2001) | 133 | 16 | 0.3 [0.0,19.4] | | |
| Gaucher (2001) | 89 | 596 | 1.7 [0.0,17.2] | • • | |
| Khang (2001) | 14 | 398 | 3.0 [0.5,12.2] | | |
| Widmark (2001) | 16 | 92 | 28.6 [2.6,1493.6] | - | • → |
| Ferrigno (2002) | 75 | 969 | 4.3 [0.1,54.8] | · | • |
| Tangerud (2002) | 7 | 78 | 3.0 [0.1,38.2] | < | |
| Attard (2003) | 42 | 351 | 2.5 [0.6,8.6] | | — |
| Friberg (2003) | 26 | 62 | 1.5 [0.2,8.3] | | |
| Tawil (2003) | 99 | 145 | 0.5 [0.0,6.1] | ← | |
| Weng (2003) | 40 | 825 | 2.4 [0.6,7.1] | | |
| Willer (2003) | 11 | 1237 | 12.2 [0.3,102.3] | · | → |
| Bergkvist (2004) | 8 | 136 | 4.6 [0.1,56.1] | < | - - |
| Preiskel (2004) | 37 | 230 | 1.6 [0.0,16.4] | · • | |
| Friberg (2005) | 11 | 380 | 9.2 [0.2,105.8] | < | → |
| Bischof (2006) | 96 | 159 | 1.7 [0.0,131.2] | · · · | |
| Jemt (2006) | 106 | 324 | 2.5 [0.5,11.8] | | |
| Total | | | 1.8 [1.3 , 2.4] | • | |
| | | | | 0.28 1.00 | 54.60 |
| | | | | favours short implants | favours long implants |

of no impact of reduced implant length on implant failure within the first year of loading could not be rejected with the exception of machined maxillary implants and rough-surfaced implants in the anterior maxilla.

There was no significant difference between narrow- and regular-diameter implants, neither for machined (OR 1.1) nor for rough surfaces (OR 1.0). Significant heterogeneity across studies was observed only in implant diameter comparison (p = 0.01). The symmetry of the funnel plot (Fig. 3) does not suggest the presence of publication bias (Sterne & Egger 2001). Descriptive analyses yielded no increased failure rates of short narrow-diameter implants compared with short regular-diameter implants in both maxillary and mandibular sites (Table 3). Failure rates of short implants supporting fixed dentures were not found to be higher than those supporting removable dentures, yet direct comparison was only possible for implants placed in the anterior mandible.

Discussion

It has been an axiom in implant dentistry that longer implants guarantee lower failure rates, although a linear relationship between implant length and success has never been proven (Lee et al. 2005). Clinical strategies to increase the surface area of short implants include the use of rough-surfaced implants as well as wider implant diameters (Tawil

Fig. 2. Forest plot: number of short (ns) and long implants (nl) in each study and odds ratios (OR) with 95% confidence intervals (CI). Studies with <10 implants in each group or OR = 0, OR = ∞ , and OR not estimable were excluded.

Table 2. Impact of implant length on implant failure rate (within-study comparison): odds ratios (OR) with 95% confidence intervals (CI_{95%}), number of studies (*N*), test-statistics for Mantel–Haenszel tests (MH), estimated failure rates after 1 year of function and sample size (*n*) of short ($\ge 7 \text{ mm}$, <10 mm) and long ($\ge 10 \text{ mm}$) implants

| | OR [CI _{95%}] | Ν | MH (<i>p</i>) | Failure rate of short implants | Failure rate of long implants |
|-------------------------|-------------------------|----|----------------------|--------------------------------|-------------------------------|
| All implants | | | | | |
| All positions | 1.8 [1.3-2.5] | 40 | $13.3 \ (p < 0.001)$ | 2.5% (n = 2223) | $1.6\% \ (n = 14, 158)$ |
| Anterior maxilla | 6.1 [2.2–17.3] | 7 | 16.7 (p < 0.001) | $4.4\% \ (n = 203)$ | $0.6\% \ (n = 801)$ |
| Posterior maxilla | 3.6 [1.4-4.9] | 13 | $12.0 \ (p = 0.001)$ | 4.1% (n = 464) | 2.3% (<i>n</i> = 1579) |
| Anterior mandible | 0.8 [0.3–2.0] | 10 | $0.4 \ (p = 0.550)$ | $1.4\% \ (n = 420)$ | 1.1% (n = 2241) |
| Posterior mandible | 0.9 [0.4–1.7] | 22 | $0.2 \ (p = 0.678)$ | $1.1\% \ (n = 934)$ | $1.7\% \ (n = 3669)$ |
| Machined-surfaced impl | ants | | - | | |
| All positions | 2.2 [1.5-3.3] | 17 | $16.2 \ (p < 0.001)$ | 4.1% (<i>n</i> = 897) | $2.2\% \ (n = 6094)$ |
| Anterior maxilla | 5.4 [1.9–15.7] | 3 | $12.8 \ (p < 0.001)$ | $6.0\% \ (n = 134)$ | $1.1\% \ (n = 440)$ |
| Posterior maxilla | 3.4 [1.7-6.6] | 4 | 15.7 (p < 0.001) | $11.8\% \ (n = 136)$ | 3.7% (<i>n</i> = 816) |
| Anterior mandible | 0.8 [0.2-3.6] | 3 | $0.1 \ (p = 0.760)$ | $1.0\% \ (n = 196)$ | $1.2\% \ (n = 1057)$ |
| Posterior mandible | 1.1 [0.5-2.3] | 8 | $0.0 \ (p = 0.856)$ | $2.2\% \ (n = 367)$ | 3.1% (<i>n</i> = 1539) |
| Rough-surfaced implants | 8 | | | | |
| All positions | 1.1 [0.6-2.1] | 22 | $0.1 \ (p = 0.733)$ | $1.2\% \ (n = 1298)$ | $0.7\% \ (n = 7544)$ |
| Anterior maxilla | ∞ | 4 | $7.9 \ (p = 0.005)$ | $1.4\% \ (n = 69)$ | $0.0\% \ (n = 361)$ |
| Posterior maxilla | 0.9 [1.7-4.2] | 9 | $0.0 \ (p = 0.826)$ | $0.9\% \ (n = 328)$ | $0.8\% \ (n = 763)$ |
| Anterior mandible | 0.7 [0.2-2.4] | 8 | $0.3 \ (p = 0.575)$ | $1.8\% \ (n = 224)$ | $1.0\% \ (n = 1184)$ |
| Posterior mandible | 0.5 [0.1–2.3] | 13 | $0.8 \ (p = 0.365)$ | $0.4\% \ (n = 555)$ | $0.7\% \ (n = 2076)$ |

& Younan 2003). In the present analysis, however, short narrow-diameter implants did not demonstrate higher failure rates when compared with their regular-diameter counterparts. The present meta-analysis suggests that roughsurfaced implants with a minimum length of 7 mm represent no risk factor for implant failure if not placed in the anterior maxilla. The clinical use of short implants in this region may, however, be limited to edentulous patients (illustrated by the small sample size of only 1% of analysed implants), as single



Fig. 3. Funnel plot for publication bias (1/standard error against log odds ratio). Only studies with finite standard error are included.

tooth replacement in the aesthetic zone generally calls for correct apico-coronal implant positioning in order to achieve long-term aesthetic results (Chen & Buser 2009).

Identification of factors impairing dental implant success is the main goal of recent research in implant dentistry (Morand & Irinakis 2007). In assessing implant failure rates, it is important to consider multiple factors together because of their interactive effects on the establishment and maintainance of osseointegration (Romeo et al. 2006). In the present meta-analysis attempts were made to control the relevant confounding variables (implant location, diameter, surface texture, timing of implant placement, loading protocol, type of denture, success criteria, and duration of follow-up). Further prosthetic factors, such as crown-to-implant ratio, splinting, occlusal table, cantilever length, implant system, opposing dentition, and bruxism, did not prove to influence short implant failure in recent investigations (Nedir et al. 2004, Tawil et al. 2006. Maló et al. 2007). One potential confounding factor is the cluster effect due to the presence of multiple implants in individual patients (Chuang & Cai 2006). Another important issue that should be kept in mind when interpreting these results is that they come from observational studies. There is controversy over the validity of evidence from non-randomized studies, as

Table 3. Descriptive analysis (data pooled across all 54 studies) for implant diameter and type of denture: failure rates (%) after 1 year of function $[CI_{95\%}]$ and sample size (*n*) of short ($\ge 7 \text{ mm}$, <10 mm) and long ($\ge 10 \text{ mm}$) implants

| | All positions | Anterior maxilla | Posterior maxilla | Anterior mandible | Posterior mandible |
|---------------------|-------------------|------------------|-------------------|-------------------|--------------------|
| Narrow diameter (| <3.75 mm) | | | | |
| Short | 1.7 [0.6-3.9] | 3.9 [0.5–13.5] | 2.0 [0.0-10.4] | 1.6 [0.2–5.5] | 0.0 [0.0-5.8] |
| | n = 293 | n = 51 | n = 51 | n = 129 | n = 62 |
| Long | 1.3 [0.9–1.8] | 1.5 [0.8–2.5] | 2.0 [0.7-4.3] | 0.9 [0.4–1.8] | 0.7 [0.1-2.6] |
| | n = 2275 | n = 919 | n = 297 | n = 785 | n = 274 |
| Regular diameter (| ≥3.75 mm) | | | | |
| Short | 2.6 [2.0-3.3] | 5.0 [2.5-8.7] | 4.7 [3.0-6.8] | 1.9 [1.0–3.4] | 1.2 [0.6–2.2] |
| | n = 2288 | n = 221 | n = 536 | n = 568 | n = 963 |
| Long | 1.7 [1.5–1.9] | 2.1 [1.5-2.7] | 2.1 [1.6–2.7] | 1.1 [0.8–1.4] | 1.9 [1.5-2.3] |
| e | n = 14,227 | n = 2416 | n = 3155 | n = 4315 | n = 4341 |
| Supporting remova | ble dentures | | | | |
| Short | 1.2 [0.2-3.5] | | | 1.2 [0.2–3.5] | |
| | n = 250 | n = 0 | n = 0 | n = 250 | n = 0 |
| Long | 1.2 [0.8–1.9] | | | 1.2 [0.8–1.9] | |
| n | <i>n</i> = 1630 | n = 0 | n = 0 | n = 1630 | n = 0 |
| Supporting fixed de | entures | | | | |
| Short | 3.2 [2.3-4.3] | 4.0 [1.6-8.1] | 5.7 [3.7-8.4] | 1.6 [0.2–5.5] | 1.2 [0.4–2.6] |
| | n = 1200 | n = 174 | n = 403 | n = 129 | n = 494 |
| Long | 2.3 [1.9-2.7] | 2.7 [1.9-3.8] | 2.8 [2.1–3.6] | 0.9 [0.4–1.6] | 2.4 [1.8-3.2] |
| | <i>n</i> = 6153 | n = 1238 | n = 1808 | n = 1231 | n = 1876 |
| All implants | 1.8 [1.6-2.0] | 2.1 [1.7-2.7] | 2.5 [2.0-3.0] | 1.1 [0.9–1.4] | 1.7 [1.4–2.1] |
| | <i>n</i> = 19,083 | n = 3607 | n = 4039 | <i>n</i> = 5797 | n = 5640 |

they are more susceptible to bias (Deeks et al. 2003). Dimensions of bias to be assessed include selection, performance, detection, attrition, and publication bias (MacLehose et al. 2000). Selection bias (concerning comparability of groups) may occur due to non-randomized treatment allocation, so particular care was given to the assessment of confounding (Higgings & Green 2008). By definition of strict inclusion criteria and quality assessment it was attempted to reduce performance bias (concerning the fidelity of the intervention), detection bias (concerning correct assessment of outcome), and attrition bias (concerning completeness of sample, follow-up and data). Quality of reporting was no issue due to the additional information provided by the authors. Performance bias may be minimized by the fact that only six studies focused on the topic of implant length. The authors attach importance to the fact that the pattern of early implant failure may not be extrapolated to long-term outcomes, although studies indicate that the first year is crucial for the success of short implants (Maló et al. 2007).

The present meta-analysis confirms the hypothesis of multicenter retrospective studies (Misch et al. 2006, Maló et al. 2007) that implant diameter increase can not compensate for length reduction. Further studies are needed in order to clearly define the limits of narrowdiameter implants with regards to clinical indications and long-term fate (Renouard & Nisand 2006). The present results are in accordance with systematic reviews that observed increased failure rates of implants shorter than 7 mm (Hagi et al. 2004, das Neves et al. 2006). A research question that still remains to be answered is from what minimal length the risk of implant failure actually increases. To compare failure rates of short implants placed in original jawbone with those of longer implants placed in the augmented bone, randomized controlled studies are needed (Graziani et al. 2004).

Acknowledgements

We would like to express special acknowledgement to all authors that contributed in this ambitious project and kindly provided additional data of their investigations: P. Åstrand, N. J. Attard, M. Bakke, P. Balleri, G. Brunel, M. Degidi, D. Deporter, A. Eliasson, N. Ferrigno, B. Friberg, A. G. Grønningsæter, T. Jemt, T. Kaus, J. T. Lambrecht, D. Lops, T. Mattsson, H. J. Meijer, P. Moheng, R. Nedir, N. Nurdin, A. Örtorp, R. Palmer, A. Piattelli, H. W. Preiskel, A. Rätzer, A. Sethi, G. Tawil, A. Visser, A. Wennerberg, D. Wismeijer.

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Supporting Information

Additional Supporting Information may be found in the online version of this article:

Table S1. Search strategy for electronic and manual searches. The number (N) of articles in the 29 screened journals are given.

Clinical Relevance

Scientific rationale for the study: Short dental implants may provide a therapeutic alternative to supplementary bone augmentation surgery in patients with limited bone volume.

Table S2. Quality assessment of the 54 included studies using the Newcastle Ottawa Scale (NOS). The NOS assesses nine criteria (category [A]-[E]): a maximum of 2 stars can be allotted in the category [E] while a maximum of 1 star can be given for all other categories. Description of star rating system: [A] studies are allotted one star in case the intervention cohort is truely or somewhat representative of the exposed individuals in the community; [B] studies are allotted one star in case the control cohort is drawn from the same community as the intervention cohort; [C] studies are allotted one star if secure records (e.g. health care records) or structured interviews are used; [D] demonstration that outcome of interest was not present at the start of the study earns one star; [E] study controls for sex and age groups earns one star, study controls for additional baseline factors earns another star; [F] studies are allotted one star in case of outcome assessment based on independent or

Principal findings: Rough-surfaced implants with a minimum length of 7 mm represent no risk factor for implant failure if not placed in the anterior maxilla.

blind assessment, reference to secure records (e.g. health records), or record linkage; [G] studies with a median duration of follow-up ≥ 6 months are allotted one star; [H] studies with a follow-up rate $\le 20\%$ are allotted one star

Appendix S1. Studies included in the meta-analysis.

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Practical implications: Bone augmentation may be obviated by the use of short rough-surfaced implants; however, it may not be indicated to compensate length reduction by implant diameter increase. This document is a scanned copy of a printed document. No warranty is given about the accuracy of the copy. Users should refer to the original published version of the material.